407 Midterm 2 Solutions¹

1. Question 1

Let X be a binomial random variable with parameters n=2 and p=1/4 so that for all integers k satisfying $0 \le k \le 2$,

$$\mathbf{P}(X=k) = {2 \choose k} p^k (1-p)^{n-k} = \frac{2!}{k!(2-k)!} (1/4)^k (3/4)^{2-k}.$$

Compute the following quantities: $\mathbf{E}X$, $\mathbf{E}(X^2)$, $\mathrm{var}(X)$

Simplify your answers to the best of your ability. (As usual, show your work.)

Solution. We have $\mathbf{P}(X=0)=(3/4)^2=9/16$, $\mathbf{P}(X=1)=2(1/4)(3/4)=6/16$, $\mathbf{P}(X=2)=(1/4)^2=1/16$, so

$$\mathbf{E}X = \sum_{x \in \mathbf{R}} x p_X(x) = 0(9/16) + 1(6/16) + 2(1/16) = 8/16 = 1/2.$$

$$\mathbf{E}X^2 = \sum_{x \in \mathbf{P}} x^2 p_X(x) = 0^2 (9/16) + 1^2 (6/16) + 2^2 (1/16) = 10/16.$$

$$var(X) = \mathbf{E}X^2 - (\mathbf{E}X)^2 = 10/16 - (1/2)^2 = (10 - 4)/16 = 6/16 = 3/8.$$

2. Question 2

Let X be a standard Gaussian random variable, so that X has PDF

$$f_X(x) = \frac{1}{\sqrt{2\pi}} e^{-x^2/2}, \quad \forall x \in \mathbf{R}.$$

Compute the following quantities: P(X = 10), P(X > 0).

Simplify your answers to the best of your ability. (As usual, show your work.) Solution. By definition of PDF,

$$\mathbf{P}(X=10) = \int_{10}^{10} f_X(x) dx = 0.$$

Also, using symmetry of the Gaussian PDF.

$$\mathbf{P}(X>0) = \int_0^\infty \frac{1}{\sqrt{2\pi}} e^{-x^2/2} dx = \frac{1}{2} \int_{-\infty}^\infty \frac{1}{\sqrt{2\pi}} e^{-x^2/2} dx = \frac{1}{2} \cdot 1 = 1/2.$$

3. Question 3

Let 0 . Let X be a geometric random variable with parameter p, so that, for any positive integer <math>k,

$$\mathbf{P}(X = k) = (1 - p)^{k - 1} p.$$

In class, we computed $\mathbf{E}X = 1/p$ and $\mathbf{E}X^2$ using a conditioning argument. For example, we showed $\mathbf{E}(X^2|X=1)=1$ and $\mathbf{E}(X^2|X>1)=1+2\mathbf{E}X+\mathbf{E}X^2$. We then solved for $\mathbf{E}X^2$ to get $\mathbf{E}X^2=(2/p^2)-1/p$.

Using this same conditioning argument (i.e. by conditioning on X = 1 and on X > 1), compute $\mathbf{E}(X^3)$. Solution. From the Total Expectation Theorem,

$$\mathbf{E}X^3 = \mathbf{E}(X^3|X=1)\mathbf{P}(X=1) + \mathbf{E}(X^3|X>1)\mathbf{P}(X>1) = \mathbf{E}(X^3|X=1)p + \mathbf{E}(X^3|X>1)(1-p).$$

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When X = 1, $X^3 = 1$ also, so $\mathbf{E}(X^3|X=1) = 1$. Also, given that X > 1, i.e. $X \neq 1$, this is equivalent to knowing that the first coin flip (defining X) is tails, and then starting over again our wait for the first heads coin flip. That is, $\mathbf{E}(X^3|X>1) = \mathbf{E}((X+1)^3) = \mathbf{E}X^3 + 3\mathbf{E}X^2 + 3\mathbf{E}X + 1$. In summary,

$$\mathbf{E}X^3 = p + (1-p)[\mathbf{E}X^3 + 3\mathbf{E}X^2 + 3\mathbf{E}X + 1].$$

That is,

$$\mathbf{E}X^{3}[1 - (1 - p)] = p + (1 - p)[3\mathbf{E}X^{2} + 3\mathbf{E}X + 1]$$

Substituting in our known values of the first and second moment,

$$\mathbf{E}X^3 = 1 + \frac{1}{p}(1-p)[3((2/p^2) - 1/p) + 3(1/p) + 1].$$

4. Question 4

Let X be binomial random variable with parameters n=2 and p=1/2. So, $\mathbf{P}(X=0)=1/4$, $\mathbf{P}(X=1)=1/2$ and $\mathbf{P}(X=2)=1/4$. And X satisfies $\mathbf{E}X=1$ and $\mathbf{E}X^2=3/2$.

Let Y be a geometric random variable with parameter 1/2. So, for any positive integer k, $\mathbf{P}(Y=k)=2^{-k}$. And Y satisfies $\mathbf{E}Y=4$ and $\mathbf{E}Y^2=32$.

Let Z be a Poisson random variable with parameter 1. So, for any nonnegative integer k, $\mathbf{P}(Z=k)=\frac{1}{e}\frac{1}{k!}$. And Z satisfies $\mathbf{E}Z=1$ and $\mathbf{E}Z^2=2$.

Let W be a discrete random variable such that $\mathbf{P}(W=-1)=1/2$ and $\mathbf{P}(W=1)=1/2$, so that $\mathbf{E}W=0$ and $\mathbf{E}W^2=1$.

Assume that X, Y and Z are all independent. Compute

$$\mathbf{E}(1 + W^{100} + W^{50}XYZ^2).$$

(You cannot assume that W is independent of X, Y, Z.)

Solution. By Definition of W, $W^{50} = W^{100} = 1$, so

$$\mathbf{E}(1 + W^{100} + W^{50}XYZ^2) = \mathbf{E}(1 + 1 + XYZ^2).$$

Since X, Y, Z are independent, we have $\mathbf{E}(XYZ^2) = \mathbf{E}X\mathbf{E}Y\mathbf{E}Z^2 = 1 \cdot 4 \cdot 2 = 8$. Finally, since the expected value of a sum is the sum of expected values, the final answer is 1 + 1 + 8 = 10.

5. Question 5

Let X_1, \ldots, X_n be independent standard Gaussian random variables. Let $Y = \min(X_1, \ldots, X_n)$ be the minimum of X_1, \ldots, X_n . What is the PDF of Y? Solution. As shown in class, the event $Y \geq t$ is equal to the event $X_1 \geq t, \ldots, X_n \geq t$. So, $\mathbf{P}(Y \geq t) = [\mathbf{P}(X_1 \geq t)]^n = [\int_t^{\infty} e^{-x^2/2} dx / \sqrt{2\pi}]^n$ for any $t \in \mathbf{R}$. We can then get the density of Y, since

$$f_Y(t) = \frac{d}{dt} \mathbf{P}(Y \le t) = \frac{d}{dt} [1 - \mathbf{P}(Y \ge t)] = -\frac{d}{dt} \left[\int_t^\infty e^{-x^2/2} dx / \sqrt{2\pi} \right]^n$$
$$= n \left(\int_t^\infty e^{-x^2/2} dx / \sqrt{2\pi} \right)^{n-1} e^{-t^2/2} / \sqrt{2\pi}.$$