

407 Midterm 2 Version 2 Solutions¹

1. QUESTION 1

TRUE/FALSE

(a) Let X be a continuous random variable with probability density function f_X . Then $f_X(x) \leq 1$ for all $x \in \mathbf{R}$.

FALSE. A density function can have value larger than 1. For example, if $f_X(x) = 2$ for any $x \in [0, 1/2]$ and $f_X(x) = 0$ otherwise, then f_X is a PDF.

(b) When X is a continuous random variable, there is a continuous function $f_X: \mathbf{R} \rightarrow [0, \infty)$ such that, for any $-\infty \leq a \leq b \leq \infty$,

$$\mathbf{P}(a \leq X \leq b) = \int_a^b f_X(x) dx$$

FALSE. f_X need not be continuous. When X is uniformly distributed in $[0, 1]$, the density f_X is discontinuous at $x = 0$ and at $x = 1$.

(c) Let X be a continuous random variable with PDF $f_X: \mathbf{R} \rightarrow [0, \infty)$. Then, for any $t \in \mathbf{R}$,

$$\frac{d}{dt} \mathbf{P}(X \leq t) = f_X(t).$$

FALSE. When X is uniformly distributed in $[0, 1]$, the density f_X is discontinuous at $x = 0$ and at $x = 1$, and $\mathbf{P}(X \leq t)$ is not differentiable at $t = 0$, so that $\frac{d}{dt} \mathbf{P}(X \leq t)$ does not exist when $t = 0$.

(d) Let X and Y be discrete random variables. Then

$$\mathbf{E}(XY) = (\mathbf{E}X)(\mathbf{E}Y).$$

FALSE. Let X be a Bernoulli random variable with parameter $0 < p < 1$, and let $Y = X$, then $\mathbf{E}XY = \mathbf{E}X^2 = \mathbf{E}X = p$ while $\mathbf{E}X\mathbf{E}Y = (\mathbf{E}X)^2 = p^2$ and $p \neq p^2$.

(e) Let A_1, \dots, A_n be disjoint events in a sample space Ω . That is, $A_i \cap A_j = \emptyset$ whenever $i, j \in \{1, \dots, n\}$ satisfy $i \neq j$. Let \mathbf{P} be a probability law on Ω . Assume $\mathbf{P}(A_i) > 0$ for all $1 \leq i \leq n$. Let $X: \Omega \rightarrow \mathbf{R}$ be a discrete random variable. Then

$$\mathbf{E}X = \sum_{i=1}^n \mathbf{P}(A_i) \mathbf{E}(X|A_i).$$

FALSE. Let $\Omega = \{1, 2, 3\}$, let \mathbf{P} be uniform on Ω , let $A_1 = \{1\}$ and let $A_2 = \{2\}$. Let X so that $X(\omega) = \omega$ for all $\omega \in \Omega$. Then $\mathbf{E}(X|A_1) = 1$ and $\mathbf{E}(X|A_2) = 2$, whereas $\mathbf{E}X = (3 + 2 + 1)/3 = 2$, but

$$\sum_{i=1}^n \mathbf{P}(A_i) \mathbf{E}(X|A_i) = (1/3)(1 + 2) = 1 \neq 2 = \mathbf{E}X.$$

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2. QUESTION 2

Let X be a discrete random variable such that

$$\mathbf{P}(X = 1) = \mathbf{P}(X = 2) = \mathbf{P}(X = 3) = 1/6, \quad \text{and}$$

$$\mathbf{P}(X = -1) = \mathbf{P}(X = -2) = \mathbf{P}(X = -3) = 1/6.$$

Compute the following quantities: $\mathbf{E}X$, $\mathbf{E}(X^2)$, $\text{var}(X)$.

Solution. By definition of X and $\mathbf{E}X$,

$$\mathbf{E}X = \sum_{x \in \mathbf{R}} xp_X(x) = (1/6)(3 + 2 + 1 - 1 - 2 - 3) = 0.$$

$$\mathbf{E}X^2 = \sum_{x \in \mathbf{R}} x^2 p_X(x) = (1/6)(3^2 + 2^2 + 1^2 + (-1)^2 + (-2)^2 + (-3)^2) = 28/6 = 14/3.$$

Lastly,

$$\text{var}(X) = \mathbf{E}X^2 - (\mathbf{E}X)^2 = 14/3 - 0 = 14/3.$$

3. QUESTION 3

Let X be an exponential random variable with parameter $\lambda = 1$, so that X has PDF

$$f_X(x) = e^{-x}, \quad \forall x \geq 0,$$

and $f_X(x) = 0$ for all $x < 0$.

Compute the following quantities: $\mathbf{E}X$, $\mathbf{P}(X > 1)$.

Solution. By definition of $\mathbf{E}X$,

$$\begin{aligned} \mathbf{E}X &= \int_{\mathbf{R}} xf_X(x)dx = \int_0^\infty xe^{-x}dx = \lim_{N \rightarrow \infty} \int_0^N x[-(d/dx)e^{-x}]dx \\ &= \lim_{N \rightarrow \infty} [-xe^{-x}]_{x=0}^{x=N} + \int_0^N e^{-x}dx = \lim_{N \rightarrow \infty} (-Ne^{-N}) + [1 - e^{-N}] = 1. \end{aligned}$$

$$\mathbf{P}(X > 1) = \int_1^\infty e^{-x}dx = \lim_{N \rightarrow \infty} \int_1^N e^{-x}dx = \lim_{N \rightarrow \infty} e^{-1} - e^{-N} = e^{-1}.$$

4. QUESTION 4

Let X and Y be discrete random variables such that $|X| \leq 10$ and $|Y| \leq 10$. Recall that $\text{cov}(X, Y) = \mathbf{E}[(X - \mathbf{E}X)(Y - \mathbf{E}Y)]$.

Prove or disprove the statement below. (In the case that you disprove the statement, it suffices to find a counterexample and explain your reasoning.)

Statement: If $\text{cov}(X, Y) = 0$, then X and Y are independent.

Solution. This statement is false. Let X be uniformly distributed in $\{-1, 0, 1\}$ and let $Y = |X|$. Then $\mathbf{E}X = (1/3)(1 + 0 + -1) = 0$, and

$$\mathbf{E}XY = \mathbf{E}X|X| = (1/3)(1 \cdot 1) + (1/3)(0) + (1/3)(-1 \cdot 1) = 0.$$

$$\mathbf{E}X\mathbf{E}Y = 0.$$

So, $\text{cov}(X, Y) = \mathbf{E}XY - \mathbf{E}X\mathbf{E}Y = 0$. However, X and Y are not independent, since

$$\mathbf{P}(X = 1, Y = 1) = \mathbf{P}(X = 1) = 1/3 \neq (1/3)(2/3) = \mathbf{P}(X = 1)\mathbf{P}(Y = 1).$$

5. QUESTION 5

Let X be a standard Gaussian random variable, so that X has PDF

$$f_X(x) = \frac{1}{\sqrt{2\pi}} e^{-x^2/2}, \quad \forall x \in \mathbf{R}.$$

Let Z be the random variable defined by

$$Z = X^4.$$

What is the PDF of Z ? (As usual, justify your answer.)

Solution. Let $t > 0$. Then

$$\mathbf{P}(Z \leq t) = \mathbf{P}(X^4 \leq t) = \mathbf{P}(X \leq t^{1/4}) = \int_{-t^{1/4}}^{t^{1/4}} \frac{1}{\sqrt{2\pi}} e^{-x^2/2} dx = 2 \int_0^{t^{1/4}} \frac{1}{\sqrt{2\pi}} e^{-x^2/2} dx.$$

So, by the Chain Rule and the Fundamental Theorem of Calculus,

$$f_Z(t) = \frac{d}{dt} \mathbf{P}(Z \leq t) = 2(1/4)t^{-3/4} \frac{1}{\sqrt{2\pi}} e^{-(t^{1/4})^2/2} = (1/2)t^{-3/4} \frac{1}{\sqrt{2\pi}} e^{-t^{1/2}/2}.$$